

On Evaluation of the Mean Service Cycle Time in Tandem Queueing Systems*

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Abstract

The problem of exact evaluation of the mean service cycle time in tandem systems of single-server queues with both infinite and finite buffers is considered. It is assumed that the interarrival and service times of customers form sequences of independent and identically distributed random variables with known mean values. We start with tandem queues with infinite buffers, and show that under the above assumptions, the mean cycle time exists. Furthermore, if the random variables which represent interarrival and service times have finite variance, the mean cycle time can be calculated as the maximum out from the mean values of these variables. Finally, obtained results are extended to evaluation of the mean cycle time in particular tandem systems with finite buffers and blocking.

Key-Words: tandem queueing systems, mean cycle time, recursive equations, independent random variables, bounds on the mean value

1 Introduction

We consider tandem systems of single-server queues with both infinite and finite buffers. The interarrival and service times of customers are assumed to form sequences of independent and identically distributed random variables. Given the mean values of interarrival and service times, we are interested in evaluating the mean cycle time of the system. In what follows, the mean cycle time is used in reference to the mean value of the time interval between

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two successive departures of customers from the system as the number of customers tends to infinity. The inverse of the mean cycle time, which implies the mean number of customers leaving the system per unit time, is referred to as system throughput.

Among other characteristics including, in particular, the mean waiting time of customers, both the mean cycle time and the throughput present system performance measures commonly used in the analysis of queues. Note that the problem of evaluating the mean waiting time is known as rather difficult; it allows for the exact solution in an explicit form only for particular queueing systems under certain restrictions on customer arrival and service processes. One can find an overview of related results in [1] (see, also, [2] for more recent results and references).

In many cases, the mean cycle time can be calculated exactly under fairly general assumptions. As an illustration, one can consider results obtained in [3, 4] in the context of investigation of stability of queueing systems. It has been shown in [4] that for a general single-server system with infinite buffer capacity, regardless of whether it is stable or not, the mean cycle time can be calculated as the maximum out from the mean values of customer interarrival and service times. Moreover, if a tandem system of queues with infinite buffers is stable, the intensities of both customer arrival and departure processes coincide, and therefore, the mean cycle time is equal to the mean interarrival time of customers.

For more complicated queueing systems including tandem queues with finite buffers and blocking, and fork-join networks there are some techniques which allow one to derive bounds on the mean cycle time. Specifically, an efficient approach which relies on results of the theory of large deviations as well as on the Perron-Frobenius spectral theory has been proposed in [5, 6]. As another example, one can consider simple bounds in [7, 8], obtained by using an approach essentially based on pure algebraic manipulations together with application of bounds on extreme values, derived in [9, 10].

In this paper, we first give quite general conditions for the mean cycle time in tandem queueing systems with infinite buffers to exist, and show how it can be calculated through the mean values of the interarrival and service times of customers. The obtained results are then extended to evaluation of the mean cycle time in particular tandem systems with finite buffers, which operate under the manufacturing and communication blocking rules.

As the starting point, we take obvious recursive equations which describe tandem system dynamics, and then examine their related explicit solution. Our approach is based on simple algebraic manipulations with the solution, combined with some classical results providing bounds on the mean value of maxima of independent and identically distributed random variables. The approach does not involve taking account of stability of the entire system, and therefore, offers a unified way of examining both stable and unstable systems.

The rest of the paper is organized as follows. In Section 2, we introduce some notations, and consider recursive equations describing the dynamics of tandem queueing systems with infinite buffers. Section 3 presents preliminary results including an existence theorem and some inequalities. Our main result which provides general existence conditions and a simple expression for calculating the mean cycle time is included in Section 4. The obtained results are then applied to the examination of the mean cycle time in particular tandem systems with finite buffers in Section 5. Finally, Section 6 offers some concluding remarks and discussion.

2 Tandem Queues with Infinite Buffers

We consider a series of M single-server queues with infinite buffers and customers of a single class. Each customer that arrives into the system is initially placed in the buffer at the 1st server and then has to pass through all the queues one after the other. Upon the completion of his service at server i , the customer is instantaneously transferred to queue $i+1$, $i = 1, \dots, M-1$, and occupies the $(i+1)$ st server provided that it is free. If the customer finds this server busy, he is placed in its buffer and has to wait until the service of all his predecessors is completed.

Denote the time between the arrivals of n^{th} customer and his predecessor by τ_{0n} , and the service time of the n^{th} customer at server i by τ_{in} , $i = 1, \dots, M$, $n = 1, 2, \dots$. Furthermore, let $D_0(n)$ be the n^{th} arrival epoch to the system, and $D_i(n)$ be the n^{th} departure epoch from the i^{th} server. We assume that for each i , $i = 0, 1, \dots, M$, the sequence $\{\tau_{in} | n = 1, 2, \dots\}$ consists of nonnegative random variables (r.v.'s).

With the condition that the tandem queueing system starts operating at time zero, and it is free of customers at the initial time, we put $D_i(0) = 0$ for all $i = 0, \dots, M$. The recursive equations representing the system dynamics can readily be written as

$$\begin{aligned} D_0(n) &= D_0(n-1) + \tau_{0n}, \\ D_m(n) &= \max(D_{m-1}(n), D_m(n-1)) + \tau_{mn}, \quad m = 1, \dots, M, \end{aligned}$$

for all $n = 1, 2, \dots$.

The above recursions can be resolved to get

$$D_m(n) = \max_{1 \leq k_1 \leq \dots \leq k_m \leq n} \left\{ \sum_{j=1}^{k_1} \tau_{0j} + \sum_{j=k_1}^{k_2} \tau_{1j} + \dots + \sum_{j=k_{m-1}}^n \tau_{mj} \right\} \quad (1)$$

for all $m = 1, \dots, M$.

We consider the evolution of the system as a sequence of service cycles: the 1st cycle starts at the initial time, and it is terminated as soon as the

M^{th} server completes its 1st service, the 2nd cycle is terminated as soon as this server completes its 2nd service, and so on. Clearly, the completion time of the n^{th} cycle can be represented as $D_M(n)$.

In many applications, one is interested in evaluating the mean service cycle time of the tandem system, which can also be treated as the mean interdeparture time of customers from the system. It is defined as

$$\gamma = \lim_{n \rightarrow \infty} \frac{1}{n} D_M(n) \quad (2)$$

provided that the above limit exists. The system throughput π presents another performance measure of interest, which is calculated as the inverse of the mean cycle time; that is, $\pi = 1/\gamma$.

3 Preliminary Results

In order to examine the existence of the mean cycle time for the tandem queueing system, we will apply the next classical theorem which has been proved in [11].

Theorem 1. *Let $\{\zeta_{ln} \mid l, n = 0, 1, \dots; l < n\}$ be a family of r.v.'s which satisfy the following properties:*

Subadditivity: $\zeta_{ln} \leq \zeta_{lk} + \zeta_{kn}$ for all $l < k < n$;

Stationarity: the joint distributions are the same for both families $\{\zeta_{ln} \mid l < n\}$ and $\{\zeta_{l+1, n+1} \mid l < n\}$;

Boundedness: for all $n = 1, 2, \dots$, there exists $\mathbb{E}[\zeta_{0n}] \geq -cn$ for some positive constant c .

Then there exists a constant γ , such that it holds

1. $\lim_{n \rightarrow \infty} \zeta_{0n}/n = \gamma$ with probability one (w.p.1),
2. $\lim_{n \rightarrow \infty} \mathbb{E}[\zeta_{0n}]/n = \gamma$.

Let us now consider some useful inequalities which will be exploited in evaluation of the mean cycle time in the next section. In what follows, we assume ξ_1, \dots, ξ_n to be independent r.v.'s.

We start with a classical result from [12], providing an upper bound on the mean value of the maximum of cumulative sums

$$\zeta_k = \xi_1 + \dots + \xi_k$$

of independent r.v.'s with zero means.

Lemma 2. *If $\mathbb{E}[\xi_k] = 0$, and $\mathbb{E}|\xi_k|^p < \infty$ for some $p > 1$, $k = 1, \dots, n$, then it holds*

$$\mathbb{E} \left[\max_{1 \leq k \leq n} |\zeta_k| \right]^p \leq 2 \left(\frac{p}{p-1} \right)^p \mathbb{E}|\zeta_n|^p.$$

The next inequality has been derived in [13]. Note that it actually remains valid under somewhat weaker conditions than that of independence between the r.v.'s ξ_1, \dots, ξ_n .

Lemma 3. *If $\mathbb{E}[\xi_k] = 0$, and $\mathbb{E}|\xi_k|^p < \infty$ for some p , $1 \leq p \leq 2$, $k = 1, \dots, n$, then it holds*

$$\mathbb{E}|\zeta_n|^p \leq \left(2 - \frac{1}{n}\right) \sum_{k=1}^n \mathbb{E}|\xi_k|^p.$$

With Lemmas 2 and 3, one can prove the following statement.

Lemma 4. *If $\mathbb{E}[\xi_k] = 0$, and $\mathbb{E}[\xi_k^2] < \infty$, $k = 1, \dots, n$, then it holds*

$$\mathbb{E} \left[\max_{1 \leq k \leq n} \zeta_k \right] \leq 2 \sqrt{\frac{2(2n-1)}{n}} \left(\sum_{k=1}^n \mathbb{E}[\xi_k^2] \right)^{1/2}.$$

Proof. First note that

$$\mathbb{E} \left[\max_{1 \leq k \leq n} \zeta_k \right] \leq \mathbb{E} \left[\max_{1 \leq k \leq n} |\zeta_k| \right] \leq \left(\mathbb{E} \left[\max_{1 \leq k \leq n} |\zeta_k| \right]^2 \right)^{1/2}.$$

By applying Lemma 2 with $p = 2$, and then Lemma 3, we get

$$\mathbb{E} \left[\max_{1 \leq k \leq n} |\zeta_k| \right]^2 \leq 8 \mathbb{E}[\zeta_n^2] \leq 8 \left(2 - \frac{1}{n} \right) \sum_{k=1}^n \mathbb{E}[\xi_k^2].$$

Finally, extracting square root leads us to the desired result. \square

Now suppose that ξ_1, \dots, ξ_n present independent and identically distributed (i.i.d.) r.v.'s. With this condition, in particular, the inequality in Lemma 4 takes the form

$$\mathbb{E} \left[\max_{1 \leq k \leq n} \zeta_k \right] \leq 2 \sqrt{2(2n-1) \mathbb{E}[\xi_1^2]}.$$

The next result obtained in [9, 10] offers an upper bound for the expected value of maximum of i.i.d. r.v.'s.

Lemma 5. *If $\mathbb{E}[\xi_1] < \infty$ and $\mathbb{D}[\xi_1] < \infty$, then it holds*

$$\mathbb{E} \left[\max_{1 \leq k \leq n} \xi_k \right] \leq \mathbb{E}[\xi_1] + \frac{n-1}{\sqrt{2n-1}} \sqrt{\mathbb{D}[\xi_1]}.$$

Assuming ξ_1, \dots, ξ_n to be i.i.d. r.v.'s, let us introduce the notation

$$\zeta_{lk} = \xi_l + \xi_{l+1} + \dots + \xi_k$$

with $1 \leq l \leq k \leq n$, and consider the following statement.

Lemma 6. If $\mathbb{E}[\xi_1] = a \leq 0$, and $\mathbb{D}[\xi_1] < \infty$, then it holds

$$\mathbb{E} \left[\max_{1 \leq l \leq k \leq n} \zeta_{lk} \right] \leq \mathbb{E}[\xi_1] + \left(4\sqrt{2(2n-1)} + \frac{n-1}{\sqrt{2n-1}} \right) \sqrt{\mathbb{D}[\xi_1]}.$$

Proof. Simple algebraic manipulations give

$$\begin{aligned} \max_{1 \leq l \leq k \leq n} \zeta_{lk} &= \max_{1 \leq l \leq k \leq n} \left\{ \sum_{i=1}^k \xi_i + \sum_{i=1}^{l-1} (-\xi_i) \right\} \\ &\leq \max_{1 \leq l \leq k \leq n} \left\{ \sum_{i=1}^k (\xi_i - a) + \sum_{i=1}^l (-\xi_i + a) + \xi_l \right\} \\ &\leq \max_{1 \leq k \leq n} \sum_{i=1}^k (\xi_i - a) + \max_{1 \leq k \leq n} \sum_{i=1}^k (-\xi_i + a) + \max_{1 \leq k \leq n} \xi_k. \end{aligned}$$

Proceeding to expectation, with $\mathbb{E}(\xi_1 - a)^2 = \mathbb{D}[\xi_1]$, we have from Lemmas 4 and 5

$$\begin{aligned} \mathbb{E} \left[\max_{1 \leq l \leq k \leq n} \zeta_{lk} \right] &\leq 4\sqrt{2(2n-1)\mathbb{D}[\xi_1]} + \mathbb{E}[\xi_1] + \frac{n-1}{\sqrt{2n-1}} \sqrt{\mathbb{D}[\xi_1]} \\ &= \mathbb{E}[\xi_1] + \left(4\sqrt{2(2n-1)} + \frac{n-1}{\sqrt{2n-1}} \right) \sqrt{\mathbb{D}[\xi_1]}. \quad \square \end{aligned}$$

4 Exact Evaluation of the Mean Cycle Time

We are now in a position to prove our main result which can be formulated as follows.

Theorem 7. Suppose that $\{\tau_{in} \mid n = 1, 2, \dots\}$, $i = 0, 1, \dots, M$, are mutually independent sequences of i.i.d. r.v.'s with $0 \leq \mathbb{E}[\tau_{i1}] < \infty$.

Then the limit at (2) exists w.p.1, and if $\mathbb{D}[\tau_{i1}] < \infty$, it is given by

$$\gamma = \max_{0 \leq i \leq M} \mathbb{E}[\tau_{i1}]. \quad (3)$$

Proof. First, we have to verify the existence of the limit at (2). In order to apply Theorem 1, let us denote

$$\zeta_{ln} = \max_{l < k_1 \leq \dots \leq k_M \leq n} \left\{ \sum_{j=l+1}^{k_1} \tau_{0j} + \sum_{j=k_1}^{k_2} \tau_{2j} + \dots + \sum_{j=k_M}^n \tau_{Mj} \right\} \quad (4)$$

for each l, n , $0 \leq l < n$, and note that we can now write

$$D_M(n) = \zeta_{0n}.$$

With simple algebraic manipulations, it is not difficult to verify that the family $\{\zeta_{ln} | l < n\}$ defined by (4) is subadditive. Since $\tau_{i1}, \tau_{i2}, \dots$, are i.i.d. r.v.'s for each $i = 0, 1, \dots, M$, the family also possesses the stationarity property. Finally, boundedness follows from the condition $0 \leq \mathbb{E}[\tau_{i1}] < \infty$ which immediately results in $\mathbb{E}[\zeta_{0n}] = \mathbb{E}[D_M(n)] \geq 0$.

Therefore, one can apply Theorem 1 so as to conclude that the limit at (2) exists w.p.1, and it can be calculated as

$$\gamma = \lim_{n \rightarrow \infty} \frac{1}{n} \mathbb{E}[D_M(n)].$$

Suppose that the maximum at (3) is achieved at some $i = m$. Consider the completion time $D_M(n)$ and represent it in the form

$$D_M(n) = \max_{1 \leq k_1 \leq \dots \leq k_M \leq n} \left\{ \sum_{j=1}^{k_1} \tau_{0j} + \sum_{j=k_1}^{k_2} \tau_{1j} + \dots + \sum_{j=k_M}^n \tau_{Mj} \right\} = \sum_{j=1}^n \tau_{mj} + \mu,$$

where

$$\begin{aligned} \mu = & \max_{1 \leq k_1 \leq \dots \leq k_M \leq n} \left\{ \sum_{j=1}^{k_1} (\tau_{0j} - \tau_{mj}) \right. \\ & + \sum_{j=k_1}^{k_2} (\tau_{1j} - \tau_{mj}) + \dots + \sum_{j=k_M}^n (\tau_{Mj} - \tau_{mj}) \\ & \left. + \tau_{mk_1} + \dots + \tau_{mk_M} \right\}. \end{aligned} \quad (5)$$

Now we can write

$$\frac{1}{n} \mathbb{E}[D_M(n)] = \mathbb{E}[\tau_{m1}] + \frac{1}{n} \mathbb{E}[\mu].$$

Let us examine the expected value $\mathbb{E}[\mu]$. With $k_1 = \dots = k_m = 1$, and $k_{m+1} = \dots = k_M = n$, we have from (5)

$$\mu \geq \tau_{01} + \tau_{11} + \dots + \tau_{m-1,1} + \tau_{m+1,n} + \dots + \tau_{Mn} \geq 0,$$

and so $\mathbb{E}[\mu] \geq 0$.

On the other hand, simple algebra gives us an obvious upper bound for μ in the form

$$\begin{aligned} \mu \leq & \max_{1 \leq k_1 \leq n} \sum_{j=1}^{k_1} (\tau_{0j} - \tau_{mj}) \\ & + \max_{1 \leq k_1 \leq k_2 \leq n} \sum_{j=k_1}^{k_2} (\tau_{1j} - \tau_{mj}) + \dots + \max_{1 \leq k_M \leq n} \sum_{j=k_M}^n (\tau_{Mj} - \tau_{mj}) \\ & + M \max(\tau_{m1}, \dots, \tau_{mn}). \end{aligned}$$

With the condition that $\mathbb{E}(\tau_{i1} - \tau_{m1}) \leq 0$ for all $i = 0, 1, \dots, M$, one can apply Lemma 6 to the first $M + 1$ terms on the right-hand side, and then Lemma 5 to the last one so as to get

$$\begin{aligned}
\mathbb{E}[\mu] &\leq \sum_{i=0}^M \left(\mathbb{E}(\tau_{i1} - \tau_{m1}) + \left(4\sqrt{2(2n-1)} + \frac{n-1}{\sqrt{2n-1}} \right) \sqrt{\mathbb{D}(\tau_{i1} - \tau_{m1})} \right) \\
&\quad + M \left(\mathbb{E}[\tau_{m1}] + \frac{n-1}{\sqrt{2n-1}} \sqrt{\mathbb{D}[\tau_{m1}]} \right) \\
&= \sum_{\substack{i=0 \\ i \neq m}}^M \mathbb{E}[\tau_{i1}] + \left(4\sqrt{2(2n-1)} + \frac{n-1}{\sqrt{2n-1}} \right) \sum_{\substack{i=0 \\ i \neq m}}^M \sqrt{\mathbb{D}(\tau_{i1} - \tau_{m1})} \\
&\quad + M \frac{n-1}{\sqrt{2n-1}} \sqrt{\mathbb{D}[\tau_{m1}]},
\end{aligned}$$

and therefore,

$$\mathbb{E}[\mu] \leq \sum_{\substack{i=0 \\ i \neq m}}^M \mathbb{E}[\tau_{i1}] + O(\sqrt{n}).$$

Finally, we have the double inequality

$$\mathbb{E}[\tau_{m1}] \leq \frac{1}{n} \mathbb{E}[D_M(n)] \leq \mathbb{E}[\tau_{m1}] + \frac{1}{n} \sum_{\substack{i=0 \\ i \neq m}}^M \mathbb{E}[\tau_{i1}] + \frac{O(\sqrt{n})}{n},$$

and with $n \rightarrow \infty$, immediately arrive at (5). \square

Corollary 8. *Under the same conditions as in Theorem 7, if at least one of the expectations $\mathbb{E}[\tau_{i1}]$, $i = 0, 1, \dots, M$, is positive, then it holds*

$$\pi = \left(\max_{0 \leq i \leq M} \mathbb{E}[\tau_{i1}] \right)^{-1}.$$

5 Tandem Queues with Finite Buffers

In this section, we show how the above approach can be applied to the analysis of tandem systems which include queues with finite buffers. Because of limited buffer capacity, servers in the systems may be blocked according to one of the blocking rules [14]. Below we present examples of systems with manufacturing blocking and communication blocking which are most commonly encountered in practice.

Let us consider a system which consists of two queues in tandem. Suppose that the buffer at the first server is infinite, while that at the second server is finite. The customers arriving to the system have to pass through the queues consecutively, and then leave the system.

First we suppose that the system operates under the manufacturing blocking rule. With this type of blocking, if upon completion of a service, the first server sees the buffer of the second one is full, it cannot be freed and has to remain busy until the second server completes its current service to provide a free space in its buffer.

For simplicity, let us assume the finite buffer to have capacity 0. With the notations introduced above, one can represent the dynamics of the system by the equations

$$\begin{aligned} D_0(n) &= D_0(n-1) + \tau_{0n}, \\ D_1(n) &= \max(\max(D_0(n), D_1(n-1)) + \tau_{1n}, D_2(n-1)), \\ D_2(n) &= \max(D_1(n), D_2(n-1)) + \tau_{2n}, \end{aligned} \quad (6)$$

for all $n = 1, 2, \dots$.

Note that from the second equation, we have $D_1(n) \geq D_2(n-1)$, and therefore, the third equation can be reduced to

$$D_2(n) = D_1(n) + \tau_{2n}.$$

Clearly, under appropriate conditions, both $\mathbb{E}[D_1(n)]/n$ and $\mathbb{E}[D_2(n)]/n$ have a common limit γ as n tends to ∞ , which can be considered as the mean cycle time of the system.

By resolving the recursive equations, we get

$$D_1(n) = \max_{1 \leq k \leq n} \left\{ \sum_{j=1}^k \tau_{0j} + \tau_{1k} + \sum_{j=k}^{n-1} \max(\tau_{1,j+1}, \tau_{2j}) \right\}.$$

As it is easy to verify, $D_1(n)$ satisfies the double inequality

$$L(n) - \max(\tau_{1,n+1}, \tau_{2n}) \leq D_1(n) \leq U(n), \quad (7)$$

where

$$\begin{aligned} L(n) &= \max_{1 \leq k \leq n} \left\{ \sum_{j=1}^k \tau_{0j} + \sum_{j=k}^n \max(\tau_{1,j+1}, \tau_{2j}) \right\}, \\ U(n) &= \max_{1 \leq k \leq n} \left\{ \sum_{j=1}^k \tau_{0j} + \sum_{j=k}^n \max(\tau_{1j}, \tau_{2,j-1}) \right\}. \end{aligned}$$

Taking into account that both $L(n)$ and $U(n)$ actually have the form of (1), one can see that, under the same conditions as in Theorem 7, it holds

$$\lim_{n \rightarrow \infty} \frac{1}{n} \mathbb{E}[L(n)] = \lim_{n \rightarrow \infty} \frac{1}{n} \mathbb{E}[U(n)] = \max(\mathbb{E}[\tau_{01}], \mathbb{E}[\max(\tau_{11}, \tau_{21})]).$$

Finally, proceeding to mean value in both sides of (7), divided by n , we conclude that the mean cycle time is given by

$$\gamma = \max(\mathbb{E}[\tau_{01}], \mathbb{E} \max(\tau_{11}, \tau_{21})).$$

Let us now assume the system to follow the communication blocking rule. This type of blocking requires the first server not to initiate service of a customer if the buffer of the second server is completed. With the finite buffer having capacity 0, the system dynamics is described by the same recursions as above, except for equation (6) which now takes the form

$$D_1(n) = \max(D_0(n), D_1(n-1), D_2(n-1)) + \tau_{1n}.$$

Resolving the recursive equations leads us to the expression

$$D_2(n) = \max_{1 \leq k \leq n} \left\{ \sum_{j=1}^k \tau_{0j} + \sum_{j=k}^n (\tau_{1j} + \tau_{2j}) \right\}.$$

Under the same conditions as in Theorem 7, we get the mean cycle time

$$\gamma = \max(\mathbb{E}[\tau_{01}], \mathbb{E}[\tau_{11}] + \mathbb{E}[\tau_{21}]).$$

6 Concluding Remarks

First note that with similar arguments as used in the proof of Theorem 7 and related lemmas, one can verify that the statement of the theorem remains valid with the condition $\mathbb{E}[\tau_{i1}^\alpha] < \infty$ for some $\alpha > 1$, instead of $\mathbb{D}[\tau_{i1}] < \infty$.

Furthermore, the proof of the theorem does not actually require that for each n , r.v.'s τ_{in} with $i = 0, 1, \dots, M$, be independent. This allows one to apply obtained results to tandem queueing systems with dependence for each customer between his interarrival and service times, including tandem queues with identical service times at each server.

Theorem 1 actually offers more general existence conditions for the mean cycle time, which imply stationarity of the sequence $\{(\tau_{0n}, \tau_{1n}, \dots, \tau_{Mn}) \mid n = 1, 2, \dots\}$ in place of independence conditions in Theorem 7.

Let us consider recursive equations describing the dynamics of the tandem queues above, and note that the symbol τ_{in} can be thought of as the n^{th} service time at server i rather than the service time of the n^{th} customer at the server. Since in this case, the order in which customers are selected from a queue for service is of no concern, the equation also describes the dynamics of systems with any queueing disciplines not permitting the pre-empting of service, and therefore, Theorem 7 can trivially be extended to such systems.

Finally note that the proof of the theorem provides us with bounds on $\mathbb{E}[D_M(n)]$, as well as an upper bound of order $n^{-1/2}$ for the convergence rate of $\mathbb{E}[D_M(n)]/n$ to the mean cycle time as n tends to ∞ .

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